

GEODA DIAGNOSTICS FOR SPATIAL REGRESSION

1. R-Square

Let the *residuals* of the spatial regression be denoted by

$$(1) \quad \hat{\varepsilon} = (\hat{\varepsilon}_1, \dots, \hat{\varepsilon}_n)' = y - \hat{y} = (y_1, \dots, y_n)' - (\hat{y}_1, \dots, \hat{y}_n)'$$

and let the *error sum of squares (ESS)* and *total sum of squares (TSS)* be given by

$$(2) \quad ESS = \hat{\varepsilon}'\hat{\varepsilon}$$

$$(3) \quad TSS = (y - \bar{y}1)'(y - \bar{y}1)$$

then in GEODA,

$$(4) \quad R\text{-square} = 1 - \frac{ESS}{TSS} \quad (= \mathbf{0.7386} \text{ for Eire example})$$

But if the *model sum of squares (MSS)* is denoted by

$$(5) \quad MSS = (\hat{y} - \bar{\hat{y}}1)'(\hat{y} - \bar{\hat{y}}1)$$

then the fraction of *TSS* that is “explained” by the model is

$$(6) \quad R^2 = \frac{MSS}{TSS} \quad (= \mathbf{0.5733} \text{ for Eire example})$$

So which do we choose?

2. Akaike Information Criterion

If the *log likelihood* of the model estimates is denoted by

$$(7) \quad L = L(\hat{\theta} | y, X) \quad (= \mathbf{-48.0591} \text{ for Eire example})$$

where $\hat{\theta} = (\hat{\theta}_1, \dots, \hat{\theta}_k)$ is the vector of parameter estimates, then the *Akaike Information Criterion (AIC)* is given by

$$(8) \quad AIC = 2 \cdot \{-L + k\} \quad (= 2 \cdot \{48.0591 + 3\} = 102.118 \text{ for Eire example in GEODA})$$

[Note: In the present case, $\hat{\theta} = (\hat{\beta}_0, \hat{\beta}_1, \hat{\sigma}^2, \hat{\lambda})$, so this *should* be $k = 4$]. Intuitively, *AIC* is an “adjusted log likelihood” that is penalized for the number of parameter fitted, in a manner analogous to *adjusted R-Square*.

3. Breusch-Pagan Test of Heteroscedasticity

The Breusch-Pagan Test considers heteroscedastic variance models of the form

$$(9) \quad \sigma_i^2 = \sigma^2(\alpha_0 + \alpha'x_i) \Rightarrow \frac{\sigma_i^2}{\sigma^2} = \alpha_0 + \alpha'x_i$$

where $x_i = (1, x_{i1}, \dots, x_{ik})$ is the vector of explanatory variables (plus intercept) for observation i and $\sigma_i^2 = \text{var}(\varepsilon_i)$. The appropriate null hypothesis is then, $H_0 : \alpha = 0$. To test this hypothesis, observe that since $\text{var}(\varepsilon_i) = E(\varepsilon_i^2)$, the residual, $\hat{\varepsilon}_i^2$, constitutes a *one-sample estimate* of σ_i^2 . If the *mean-square error* is denoted by

$$(9) \quad s^2 = \frac{1}{n} \hat{\varepsilon}'\hat{\varepsilon} = \frac{1}{n} ESS$$

then under H_0 , the sample vector

$$(10) \quad r = \left(\frac{\hat{\varepsilon}_1^2}{s^2}, \dots, \frac{\hat{\varepsilon}_n^2}{s^2} \right)$$

is a natural estimate of $(\sigma_1^2 / \sigma^2, \dots, \sigma_n^2 / \sigma^2)$. Hence if one regresses r on the set of explanatory variables, $X = [1, x_1, \dots, x_k]$, then “significantly large” values for the *model sum of squares* (*MSS*) of this regression (under the null hypothesis H_0) indicate that model (9) fits better than would be expected under H_0 . The appropriate *Breusch-Pagan statistic*, S_{BP} , is thus taken to be twice this *MSS*,

$$(11) \quad S_{BP} = \frac{1}{2} MSS = \frac{1}{2} [r'X(XX)^{-1}Xr - n] \quad (= 0.0743 \text{ for Eire example})$$

which can be shown to be asymptotically distributed χ_{k+1}^2 under H_0 . In the Eire case, this value is not sufficiently high to suggest the presence of significant heteroscedasticity (*P-value* = .785)